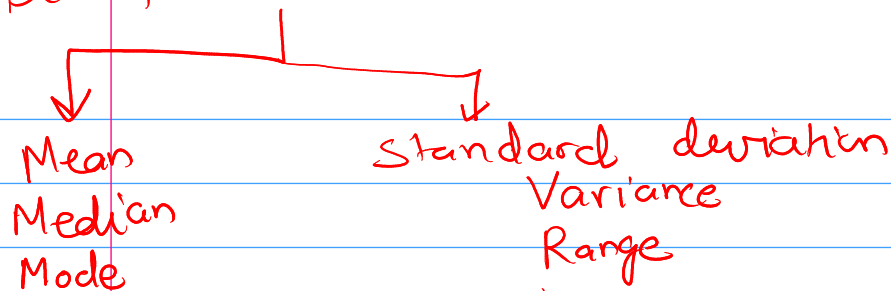


Descriptive Statistics



Inferential Statistics

Probability.



- Laws of Probability → Conditional Prob.
- Random Variables — Bayes' Theorem
- Probability Mass function
- Counting the number of events.

Laws of Probability:

Probability theory studies assignments of prob. to "Events".

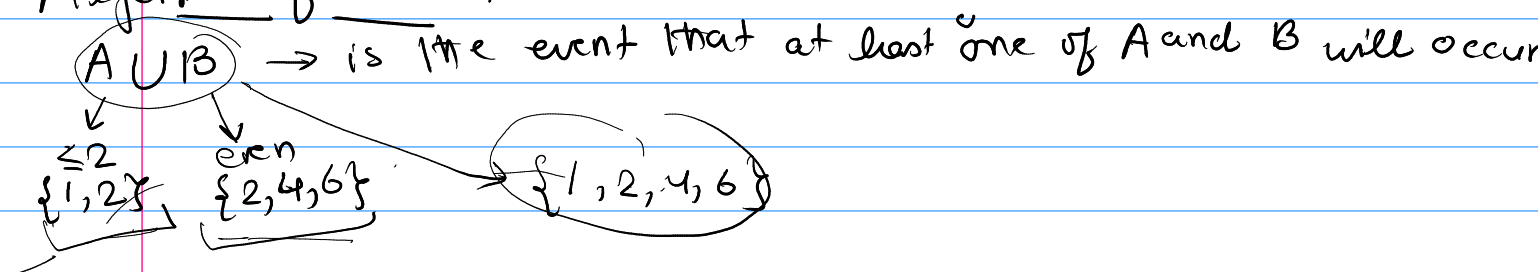
Events are statements like propositions for which it is not immediately clear whether they are true or false. eg:

↳ It will rain in two weeks from now.

↳ India's population will take over China's by 2022.

↳ '3' shows up when you roll a fair die.

Algebra of Events:



$A \cap B$ → is the event that both A and B will occur

↳ {2}

A^c → is the event that A will not occur.

↳ {3, 4, 5, 6}

We also define an event Ω which is the union of all possible events; \emptyset is the intersection of all possible events.

$$A \cap A^c = \emptyset \rightarrow \text{Empty set}$$

$$\Omega^c = \emptyset$$

$$A \cup A^c = \Omega$$

Rolling of a fair die.

$A = \{1, 2\}$ Numbers less than or equal to 2 show up.

$B = \{2, 4, 6\}$ Even Numbers show up.

$$P(A) = 1/3 = 1/6 + 1/6$$

$$P(B) = 1/2$$

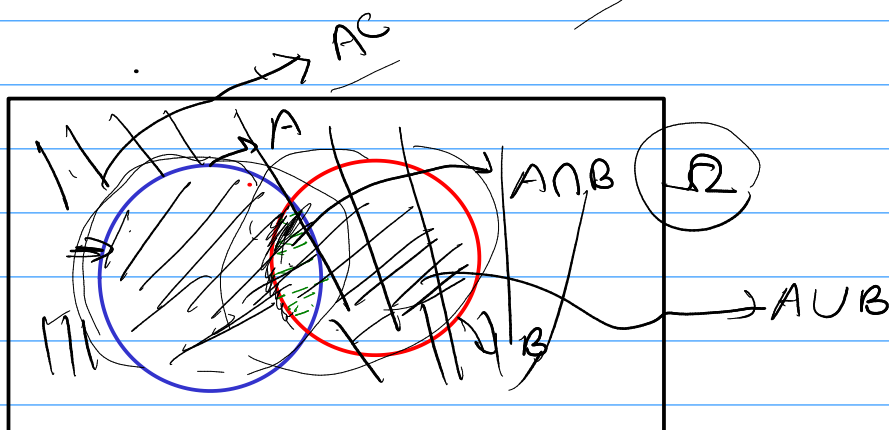
$$P(A \cap B) = P(\{2\}) = 1/6$$

$$P(A \cup B) = P(\{1, 2, 4, 6\}) = 2/3$$

$$P(A) + P(B) = \frac{5}{6}$$

$$- P(A \cap B) = \frac{2}{3} = P(A \cup B)$$

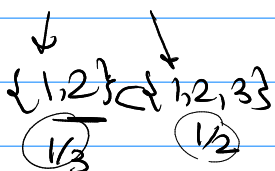
$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$



Laws of Probability:

$P(A)$ is a real number b/w 0 and 1, the prob. of (occurrence of) the event A.

- Prob. of the universal event is 1, i.e., $P(\Omega) = 1$
- If $A \subset B$, then $P(A) \leq P(B)$

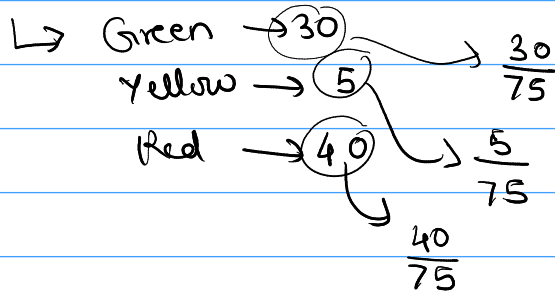


• The "excluded middle" law, $P(A) + P(A^c) = 1$

• The addition law:

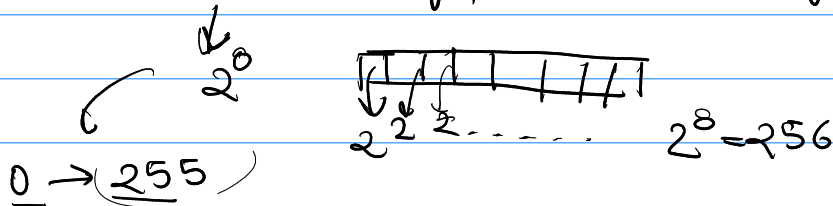
$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

Ex: Traffic lights at an intersection cycle through every 75 seconds.

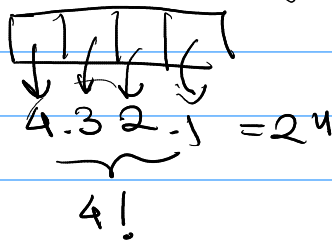


• Counting the sizes of various sets.

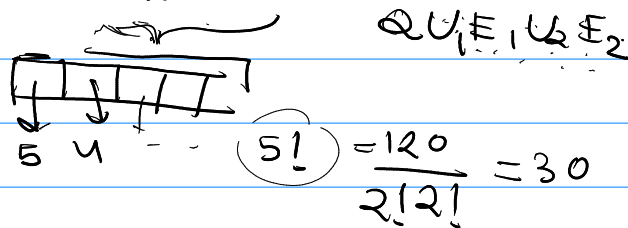
Ex: Find the total number of possible 8-bit bytes: 000 101 00



Ex: Find the # of four letter sequences that can be obtained by ordering the letters A, B, C, D without repetitions; A D C B



Ex: Find the number of unique ^{five-letter} sequences that can be formed from the word QUEUE?



Ex: $0! = 1$? $5! = 5 \times 4 \times 3 \times 2 \times 1$

$$n! = n \cdot (n-1)!$$

$$= n \cdot (n-1) \cdot (n-2) \cdot \dots \cdot 1 = n(n-1)!$$

$$(n-1)! = \frac{n!}{n}$$

\leftarrow
 $n=1$

$$\boxed{0! = \frac{1!}{1} = 1!}$$

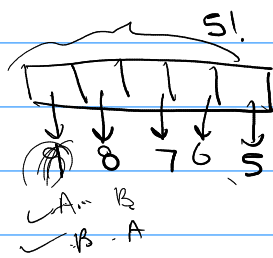
$$n=0 \quad (-1)! = \frac{0!}{0} \rightarrow \text{undefined}$$

\hookrightarrow We do not define factorials for negative numbers.

Ex: Suppose nine basketball players are labeled by the letters A, B, C, D, E, F, G, H, I

A lineup consists of five players. The order of the players in the lineup is not important

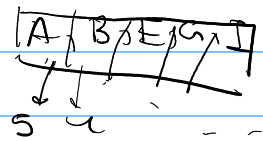
\hookrightarrow # of distinct lineups?



$$9 \cdot 8 \cdot 7 \cdot 6 \cdot 5 = \frac{9!}{4!}$$

Had the order been important

$$\rightarrow {}^9P_5 = \frac{9!}{(9-5)!}$$



$$\frac{9!}{4!5!} = 126$$

$${}^9C_5 = {}^9C_4$$

$$\binom{9}{5} = {}^9C_5$$

$$\boxed{{}^nC_k = \frac{n!}{k!(n-k)!} = {}^nC_{n-k}}$$

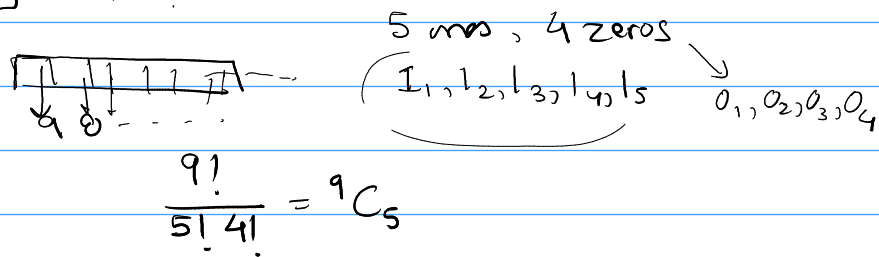
$$\boxed{{}^nP_r = \frac{n!}{(n-r)!}}$$

\hookrightarrow when order is important

\hookrightarrow order is not important

Ex: How many binary seqⁿ of length 9 have exactly 5 ones?

$$\rightarrow \binom{9}{5} = \frac{9!}{5!4!} = 126$$



Ex: Suppose there are 9 socks loose in a drawer in a dark room

6 orange 3 Blue

Someone selects two socks at random. What is the prob^o that the two socks are of the same color?

$$\rightarrow |\Omega| = \binom{9}{2} = {}^9C_2 = \frac{9 \cdot 8}{2} = 36$$

$$\text{When socks are orange } \binom{6}{2} = {}^6C_2 = \frac{6 \cdot 5}{2} = 15$$

$$\text{When socks are blue } \binom{3}{2} = 3$$

$$P(\text{socks are of same color}) = \frac{15 + 3}{36} = \frac{1}{2}$$

Alternative Method:

$$\rightarrow \text{Orange } \frac{6}{9} = \frac{2}{3}$$

$$\rightarrow \text{Blue } \frac{3}{9} = \frac{1}{3}$$

A is that orange colored sock is picked at first

Sock 1 is orange $\left(\frac{6}{9}\right)$ and Sock 2 is also orange $\left(\frac{5}{8}\right) = \frac{30}{72}$ is the event that we pick orange colored sock in the 2nd draw

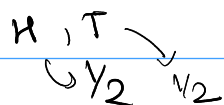
sock 1 is Blue and sock 2 is also blue

$$\frac{3}{9} \cdot \frac{2}{8} = \frac{6}{72} \rightarrow \frac{30}{72} + \frac{6}{72} = \frac{36}{72} = \frac{1}{2}$$

* Conditional Prob. and Independence

Two events A and B are said to be independent if

$$P(A \cap B) = P(A) \cdot P(B)$$



$$P(\{HH\}) = \frac{1}{2} \cdot \frac{1}{2}$$

A_1, A_2, \dots, A_n

$$P(A_1 \cap A_2 \dots \cap A_n) = P(A_1) P(A_2) \dots P(A_n)$$

"Mutual Independence" A_1, A_2, A_3

$$\left\{ \begin{array}{l} \perp P(A_1 \cap A_2 \cap A_3) = P(A_1) P(A_2) P(A_3) \\ \perp P(A_1 \cap A_2) = P(A_1) \cdot P(A_2) \\ \perp P(A_1 \cap A_3) = P(A_1) \cdot P(A_3) \\ \perp P(A_2 \cap A_3) = P(A_2) P(A_3) \end{array} \right.$$

A_1, A_2, A_3, A_4

$$P(A_1 \cap A_2 \cap A_3) = P(A_1) P(A_2) P(A_3)$$

• Mutually Exclusive Events:

A and B are mutually exclusive iff $A \cap B = \emptyset$

$A, B = A^c$, A and A^c are M-E.
 Ω and \emptyset are M-E.

$$P(A \cup B) = P(A) + P(B)$$

- Events can also be dependent sometimes, meaning the occurrence of one event is likely to have a bearing on the other event.

↳ Conditional Prob. → A, B with $P(B) \neq 0$

$$P(A|B) \neq P(A)$$

conditional prob. of A given B

"Knowing that B is true, what is the prob. of occurrence of A"

g) $P(B) \neq 0$,

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{n(A \cap B) / n(\Omega)}{n(B) / n(\Omega)} = \frac{n(A \cap B)}{n(B)}$$

A: odd numbers show up
 B: even numbers show up
 $\frac{1}{2}$ $P(A)$
 $\frac{1}{2}$ $P(B)$

A: numbers ≤ 2 show up $\rightarrow \{1, 2\}$

B: even numbers ≤ 4 show up $\rightarrow \{2, 4\}$
 50% 50%

$$P(A) = \frac{1}{3}$$

$P(A|\Omega)$

$P(A \cap B) = 0$

$P(A|B) = \frac{1}{2}$

$P(A|B) = 0$

$A \cap B = \{2\}$

$n(B) = 2$

$n(A \cap B) = 1$

$$P(A|B) = \frac{1}{2}$$

• If two events are mutually exclusive, then $P(A|B) = 0$

• If two events are independent, $P(A \cap B) = P(A) \cdot P(B)$

$$P(A|B) = P(A)$$

* Bayes' Theorem: $P(A) \neq 0, P(B) \neq 0$

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

$$P(B|A) = \frac{P(A \cap B)}{P(A)} \rightarrow P(A \cap B) = P(B|A) \cdot P(A)$$

$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B)}$$

$\{x, y, z\} \rightarrow \{y, z, x\}$ $P(\text{rel.} | \text{prob. measure})$

Ex: What is the prob. that one of the dice shows 2 when the sum is 7. (we roll two fair dice)

$n(B) = (1, 6), (2, 5), (3, 4), (4, 3), (5, 2), (6, 1)$

$n(A \cap B) = 2$

$P(A|B) = \frac{2}{6} = \frac{1}{3}$

$$A = \{ (2,1), (2,2), (2,3), \dots, (2,6) \}$$

$$\{ (1,2), (3,2), (4,2), (5,2), (6,2) \}$$

$$n(A) = 11$$

$$P(A) = 11/36$$

$$P(A|B) = 1/3$$

* Non-intuitive example:

$$\rightarrow \frac{1}{1000} = 0.001$$

For every 1000 persons being tested, only one traveler is expected to be infected. The RT-PCR test has a false positive in 5% cases, in which the traveler is healthy & uninfected.

But the test is highly sensitive, and never fails to detect a truly infected person. X returned with a positive infection in the test. What is the prob. that X was indeed infected?

$$\frac{1}{50}$$

$$\rightarrow 0.95$$

$$P(\text{infected} | \text{pos}) = \frac{P(\text{infected and pos})}{P(\text{pos})}$$

$$P(\text{infected}) = \frac{1}{1000} = 0.001$$

$$= \frac{P(\text{pos} | \text{infected}) P(\text{infected})}{P(\text{pos})} \quad \text{Bayes' Theorem}$$

$$= \frac{0.001}{0.05095} = \frac{0.001}{0.05095} = 0.019627 \approx 0.02$$

$$\rightarrow P(\text{pos}) = P(\text{pos} | \text{infected}) P(\text{infected}) + P(\text{pos} | \text{uninfected}) P(\text{uninfected})$$

$$= P(\text{pos} \cap \text{infected}) + P(\text{pos} \cap \text{uninfected})$$

$$= 0.001 + 0.05 * 0.999$$

$$= 0.05095$$

We can also arrive at the same solⁿ using a different approach.

999 healthy 1 infected

$$\# \text{ of positive tests} = 999 \times 0.05 + 1$$

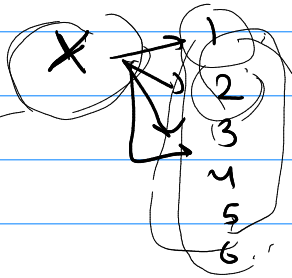
$$= 50.95$$

$$\therefore \text{Prob. of choosing an infected person} = \frac{1}{50.95} \approx 0.02$$

Random Variables

A real valued function on Ω .

$$X_1 \rightarrow X_2$$



Roll two fair dice

$S \rightarrow$ sum of numbers showing up

$$S = 2, 3, 4, \dots \quad S = X_1 + X_2$$

H, T
 $X = \begin{cases} 0 & \text{when H} \\ 1 & \text{when T} \end{cases}$

(1,1) (1,2), (1,3), ...
 2, 3, ...

$$\bar{X} = \frac{1+2+3+4+5+6}{6} = 3.5 \text{ rand}$$

$X \rightarrow$ Discrete-type random variable

$$X \in \{0, 1\}$$

$$X \in \{1, 2, 3, 4, 5, 6\}$$

$$X \in \{1, 2, 3, \dots\}$$

$$\{1, \sqrt{2}, \pi, \dots\}$$

Discrete type random var.

Prob. Mass. function

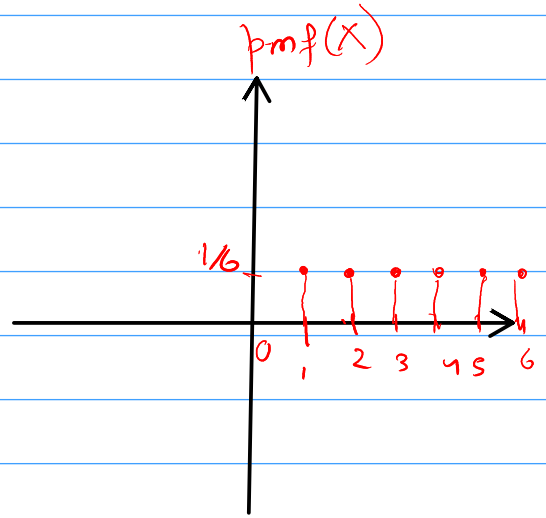
Population Mass " (pmf)

u ∈ U, ...

$$P(X = u)$$

{1, 2, 3, 4, 5, 6}

$$\begin{cases} P(X=1) = 1/6 \\ P(X=2) = 1/6 \\ \vdots \\ P(X=6) = 1/6 \\ P(X=7) = 0 \end{cases}$$



Ex Toss a fair coin twice.
 ↗ H X=0
 ↘ T X=1

Sum of two independent coin tosses.
 ↙ H ↘ T

$$P(S=0) = \frac{1}{2} \cdot \frac{1}{2} = 0.25$$

$$P(S=1) = \frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2} = 0.5$$

$$P(S=2) = \frac{1}{2} \cdot \frac{1}{2} = 0.25$$

